



Regularized PCA to denoise and visualise data

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Outline

1 PCA

2 Regularized PCA

MSE point of view

Bayesian point of view

3 Results



Geometrical point of view

Maximize the variance of the projected points

Minimize the reconstruction error

⇒ Approximation of $\mathbf{X}$ of low rank ($S < p$)

$$\|\mathbf{X}_{n \times p} - \hat{\mathbf{X}}_{n \times p}\|^2 \quad \text{SVD: } \hat{\mathbf{X}}^{PCA} = \mathbf{U}_{n \times S} \mathbf{\Lambda}_{S \times S}^{\frac{1}{2}} \mathbf{V}'_{p \times S}$$

$\mathbf{F} = \mathbf{U} \mathbf{\Lambda}^{\frac{1}{2}}$ principal components

$\mathbf{V}$ principal axes

Alternating least squares algorithm: $\|\mathbf{X} - \mathbf{FV}'\|^2$

$$\mathbf{F} = \mathbf{XV}(\mathbf{V}'\mathbf{V})^{-1}$$

$$\mathbf{V} = \mathbf{X}'\mathbf{F}(\mathbf{F}'\mathbf{F})^{-1}$$



Model in PCA

$$x_{ij} = \text{signal} + \text{noise}$$

- Random effect model: Factor Analysis, Probabilistic PCA (Lawley, 1953, Roweis, 1998, Tipping & Bishop, 1999)
 - ⇒ individuals are exchangeable
 - ⇒ relationship between variables
- Fixed effect model (Lawley, 1941, Caussinus 1986)
 - ⇒ individuals are not i.i.d
 - ⇒ study the individuals and the variables
 - sensory analysis (products - sensory descriptors)
 - plant breeding (genotypes - environments)



Fixed effect model

$$\mathbf{X}_{n \times p} = \tilde{\mathbf{X}}_{n \times p} + \varepsilon_{n \times p}$$

$$x_{ij} = \sum_{s=1}^S \sqrt{d_s} q_{is} r_{js} + \varepsilon_{ij}, \quad \varepsilon_{ij} \sim \mathcal{N}(0, \sigma^2)$$

- Randomness due to the noise
- Individuals have different expectations: individual parameters
- Inferential framework: when $n \nearrow$ the number of parameters $\nearrow$
 $\Rightarrow$ Asymptotic: $\sigma \rightarrow 0$
- Maximum likelihood estimates ($\hat{\theta}$) = least square estimates
 EM algorithm = ALS algorithm
 $\Rightarrow$ Not necessarily the best in MSE

$\Rightarrow$ **Shrinkage estimates**: minimizes $\mathbb{E}(\phi \hat{\theta} - \theta)^2$



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Minimizing the MSE

$$x_{ij} = \tilde{x}_{ij} + \varepsilon_{ij} = \sum_{s=1}^S \tilde{x}_{ij}^s + \varepsilon_{ij}$$

$$\text{MSE} = \mathbb{E} \left(\sum_{i,j} \left(\sum_{s=1}^{\min(n-1,p)} \phi_s \hat{x}_{ij}^{(s)} - \tilde{x}_{ij} \right)^2 \right) \text{ with } \hat{x}_{ij}^{(s)} = \sqrt{\lambda_s} u_{is} v_{js}$$

$$\text{MSE} = \mathbb{E} \left(\sum_{i,j} \left(\sum_{s=1}^S \phi_s \hat{x}_{ij}^{(s)} - \tilde{x}_{ij}^{(s)} \right)^2 + \sum_{s=S+1}^{\min(n-1,p)} \left(\phi_s \hat{x}_{ij}^{(s)} - \tilde{x}_{ij}^{(s)} \right)^2 \right)$$

For all $s \geq S+1$, $\tilde{x}_{ij}^{(s)} = 0$, therefore $\phi_{S+1} = \dots = \phi_{\min(n-1,p)} = 0$.



Minimizing the MSE

$$\text{MSE} = \mathbb{E} \left(\sum_{i,j} \left(\sum_{s=1}^S \phi_s \hat{x}_{ij}^{(s)} - \tilde{x}_{ij} \right)^2 \right) \text{ with } \hat{x}_{ij}^{(s)} = \sqrt{\lambda_s} u_{is} v_{js}$$

$$\phi_s = \frac{\sum_{i,j} \mathbb{E}(\hat{x}_{ij}^{(s)}) \tilde{x}_{ij}}{\sum_{i,j} \mathbb{E}(\hat{x}_{ij}^{(s)2})} = \frac{\sum_{i,j} \mathbb{E}(\hat{x}_{ij}^{(s)}) \tilde{x}_{ij}}{\sum_{i,j} \left(V(\hat{x}_{ij}^{(s)}) + \left(\mathbb{E}(\hat{x}_{ij}^{(s)}) \right)^2 \right)}$$

When $\sigma \rightarrow 0$: $\mathbb{E}(\hat{x}_{ij}^{(s)}) = \tilde{x}_{ij}^{(s)}$ (unbiased); $V(\hat{x}_{ij}^{(s)}) = \frac{\sigma^2}{\min(n-1; p)}$
(Pazman & Denis, 1996; Denis & Gower, 1999)

$$\phi_s = \frac{\sum_{i,j} \tilde{x}_{ij}^{(s)2}}{\sum_{i,j} \frac{\sigma^2}{\min(n-1; p)} + \sum_{i,j} \tilde{x}_{ij}^{(s)2}}$$



Minimizing the MSE

$$x_{ij} = \tilde{x}_{ij} + \varepsilon_{ij} = \sum_{s=1}^S \sqrt{d_s} q_{is} r_{js} + \varepsilon_{ij}, \quad \varepsilon_{ij} \sim \mathcal{N}(0, \sigma^2)$$

$$\text{MSE} = \mathbb{E} \left(\sum_{i,j} \left(\sum_{s=1}^S \phi_s \hat{x}_{ij}^{(s)} - \tilde{x}_{ij} \right)^2 \right) \quad \text{with} \quad \hat{x}_{ij}^{(s)} = \sqrt{\lambda_s} u_{is} v_{js}$$

$$\begin{aligned} \phi_s &= \frac{\sum_{i,j} \tilde{x}_{ij}^{(s)^2}}{\sum_{i,j} \frac{\sigma^2}{\min(n-1;p)} + \sum_{i,j} \tilde{x}_{ij}^{(s)^2}} = \frac{d_s}{\frac{np\sigma^2}{\min(n-1;p)} + d_s} \\ &= \frac{\text{variance signal}}{\text{variance total}} \end{aligned}$$

$$\hat{\phi}_s = \frac{\lambda_s - \hat{\sigma}^2}{\lambda_s} \quad \text{for } s = 1, \dots, S$$



Regularized PCA

$$\hat{x}_{ij}^{\text{PCA}} = \sum_{s=1}^S \sqrt{\lambda_s} u_{is} v_{js}$$

$$\hat{x}_{ij}^{\text{rPCA}} = \sum_{s=1}^S \left(\frac{\lambda_s - \hat{\sigma}^2}{\lambda_s} \right) \sqrt{\lambda_s} u_{is} v_{js} = \sum_{s=1}^S \left(\sqrt{\lambda_s} - \frac{\hat{\sigma}^2}{\sqrt{\lambda_s}} \right) u_{is} v_{js}$$

Compromise between hard and soft thresholding

σ^2 small $\rightarrow$ regularized PCA $\approx$ PCA

σ^2 large $\rightarrow$ individuals coordinates toward the center of gravity

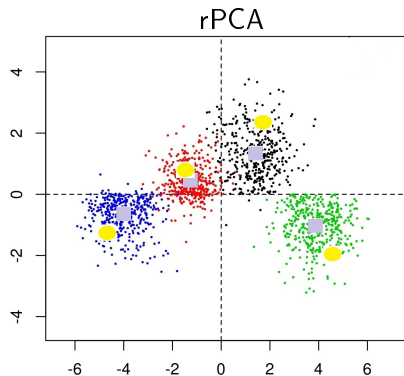
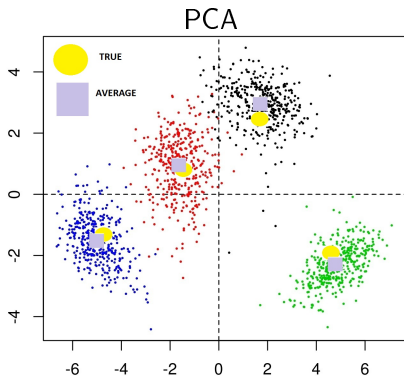
$$\hat{\sigma}^2 = \frac{RSS}{ddl} = \frac{n \sum_{s=S+1}^q \lambda_s}{np - p - nS - pS + S^2 + S} \quad (\mathbf{X}_{n \times p}; \mathbf{U}_{n \times S}; \mathbf{V}_{p \times S})$$

(only asympt unbiased)

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Illustration of the regularization

$$\mathbf{X}_{4 \times 10}^{sim} = \tilde{\mathbf{X}}_{4 \times 10} + \varepsilon_{4 \times 10}^{sim} \quad sim = 1, \dots, 1000$$



⇒ rPCA more biased less variable



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$$Y = \mathbf{X}\beta + \varepsilon, \varepsilon \sim \mathcal{N}(0, \sigma^2)$$

$$\beta \sim \mathcal{N}(0, \tau^2)$$

$$\hat{\beta}^{\text{MAP}} = (\mathbf{X}'\mathbf{X} + \kappa)^{-1}\mathbf{X}'Y, \text{ with } \kappa = \sigma^2/\tau^2$$

- ③ Results



Probabilistic PCA (*Roweis, 1998, Tipping & Bishop, 1999*)

⇒ A specific case of an exploratory factor analysis model

$$\mathbf{x}_{ij} = (\mathbf{Z}_{n \times S} \mathbf{B}'_{p \times S})_{ij} + \varepsilon_{ij}, \mathbf{z}_i \sim \mathcal{N}(0, \mathbb{I}_S), \varepsilon_i \sim \mathcal{N}(0, \sigma^2 \mathbb{I}_p)$$

- Conditional independence : $x_i | z_i \sim \mathcal{N}(\mathbf{B}z_i, \mathbb{I}_p)$
- Distribution for the observations:
 $\mathbf{x}_i \sim \mathcal{N}(0, \Sigma)$ with $\Sigma = \mathbf{B}\mathbf{B}' + \sigma^2 \mathbb{I}_p$
- Explicit solution:
 - $\hat{\sigma}^2 = \frac{1}{p-S} \sum_{s=S+1}^p \lambda_s$
 - $\hat{\mathbf{B}} = \mathbf{V}(\Lambda - \sigma^2 \mathbb{I}_S)^{\frac{1}{2}}$



Probabilistic PCA

$$\mathbf{x}_{ij} = (\mathbf{Z}_{n \times S} \mathbf{B}'_{p \times S})_{ij} + \varepsilon_{ij}, \mathbf{z}_i \sim \mathcal{N}(0, \mathbb{I}_S), \varepsilon_i \sim \mathcal{N}(0, \sigma^2 \mathbb{I}_p)$$

$\Rightarrow$ Random effect model: no individual parameters

$\Rightarrow$ BLUP $\mathbb{E}(\mathbf{z}_i | \mathbf{x}_i)$: $\hat{\mathbf{Z}} = \mathbf{X} \hat{\mathbf{B}} (\hat{\mathbf{B}}' \hat{\mathbf{B}} + \sigma^2 \mathbb{I}_S)^{-1}$

$$\hat{\mathbf{X}}^{\text{pPCA}} = \hat{\mathbf{Z}} \hat{\mathbf{B}}'$$

Using $\hat{\mathbf{B}} = \mathbf{V}(\Lambda - \sigma^2 \mathbb{I}_S)^{\frac{1}{2}}$ and SVD $\mathbf{X} = \mathbf{U} \Lambda^{1/2} \mathbf{V}' \rightarrow \mathbf{X} \mathbf{V} = \Lambda^{1/2} \mathbf{U}$

$$\hat{\mathbf{X}}^{\text{pPCA}} = \mathbf{U}(\Lambda - \sigma^2 \mathbb{I}_S) \Lambda^{-\frac{1}{2}} \mathbf{V} \quad \hat{x}_{ij}^{\text{pPCA}} = \sum_{s=1}^S \left(\sqrt{\lambda_s} - \frac{\hat{\sigma}^2}{\sqrt{\lambda_s}} \right) u_{is} v_{js}$$

Bayesian treatment of the fixed effect model with an a priori on $\mathbf{z}_i$:

$\Rightarrow$ Regularized PCA



Probabilistic PCA

⇒ EM algorithm in pPCA: two ridge regressions

$$\text{Step E:} \quad \hat{\mathbf{Z}} = \mathbf{X}\hat{\mathbf{B}}(\hat{\mathbf{B}}'\hat{\mathbf{B}} + \hat{\sigma}^2\mathbb{I}_S)^{-1}$$

$$\text{Step M:} \quad \hat{\mathbf{B}} = \mathbf{X}'\hat{\mathbf{Z}}(\hat{\mathbf{Z}}'\hat{\mathbf{Z}} + \hat{\sigma}^2\Lambda^{-1})^{-1}$$

⇒ EM algorithm in PCA: two regressions $\|\mathbf{X} - \mathbf{F}\mathbf{V}'\|^2$

$$\mathbf{F} = \mathbf{X}\mathbf{V}(\mathbf{V}'\mathbf{V})^{-1}$$

$$\mathbf{V} = \mathbf{X}'\mathbf{F}(\mathbf{F}'\mathbf{F})^{-1}$$



An empirical Bayesian approach

$$\text{Model: } x_{ij} = \sum_{s=1}^S \tilde{x}_{ij}^{(s)} + \varepsilon_{ij}, \quad \varepsilon_{ij} \sim \mathcal{N}(0, \sigma^2)$$

$$\text{Prior: } \tilde{x}_{ij}^{(s)} \sim \mathcal{N}(0, \tau_s^2)$$

$$\text{Posterior: } \mathbb{E} \left(\tilde{x}_{ij}^{(s)} | x_{ij}^{(s)} \right) = \Phi_s x_{ij}^{(s)} \text{ with } \Phi_s = \frac{\tau_s^2}{\tau_s^2 + \frac{1}{\min(n-1; p)} \sigma^2}$$

Maximizing the likelihood of $\left(x_{ij}^{(s)} \right) \sim \mathcal{N}(0, \tau_s^2 + \frac{1}{\min(n-1; p)} \sigma^2)$ as a function of τ_s^2 to obtain: $\hat{\tau}_s^2 = \left(\frac{1}{np} \lambda_s - \frac{1}{\min(n-1; p)} \hat{\sigma}^2 \right)$

$$\hat{\Phi}_s = \frac{\lambda_s - \frac{np}{\min(n-1; p)} \hat{\sigma}^2}{\lambda_s}$$

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Simulation design

The simulated data:

$$\mathbf{X} = \tilde{\mathbf{X}} + \varepsilon$$

Vary several parameters to construct $\tilde{\mathbf{X}}$:

- n/p : $100/20 = 5$, $50/50 = 1$ and $20/100 = 0.2$
- S underlying dimensions: 2, 4, 10
- the ratio d_1/d_2 : 4, 1

ε_{ij} is drawn from a Gaussian distribution with a standard deviation chosen to obtain a signal-to-noise ratio (SNR): 4, 1, 0.8



Soft thresholding and SURE (Candès *et al.*, 2012)

$$\hat{x}_{ij}^{\text{Soft}} = \sum_{s=1}^{\min(n,p)} \left(\sqrt{\lambda_s} - \lambda \right)_+ u_{is} v_{js} \quad \|\mathbf{X} - \tilde{\mathbf{X}}\|^2 + \lambda \|\tilde{\mathbf{X}}\|_*$$

Choosing the tuning parameter λ ? $\text{MSE}(\lambda) = \mathbb{E} \|\tilde{\mathbf{X}} - \text{SVT}_\lambda(\mathbf{X})\|^2$

$\Rightarrow$ Stein's Unbiased Risk Estimate (SURE)

$$\text{SURE}(\text{SVT}_\lambda)(\mathbf{X}) = -np\sigma^2 + \sum_i^{\min(n,p)} \min(\lambda^2, \lambda_i) + 2\sigma^2 \text{div}(\text{SVT}_\lambda)(\mathbf{X})$$

$\Rightarrow \lambda$ is obtained minimizing SURE. Remark: $\hat{\sigma}^2$ is required

Simulation study

n/p	S	SNR	d_1/d_2	$MSE(\hat{\mathbf{X}}^{\text{PCA}}, \tilde{\mathbf{X}})$	$MSE(\hat{\mathbf{X}}^{\text{rPCA}}, \tilde{\mathbf{X}})$	$MSE(\hat{\mathbf{X}}^{\text{SURE}}, \tilde{\mathbf{X}})$
100/20	4	4	4	8.25E-04	8.24E-04	1.42E-03
100/20	4	4	1	8.26E-04	8.25E-04	1.43E-03
100/20	4	1	4	2.60E-01	1.96E-01	2.43E-01
100/20	4	1	1	2.47E-01	2.04E-01	2.62E-01
100/20	4	0.8	4	7.41E-01	4.27E-01	4.36E-01
100/20	4	0.8	1	6.68E-01	4.40E-01	4.83E-01
50/50	4	4	4	5.48E-04	5.48E-04	1.04E-03
50/50	4	4	1	5.46E-04	5.46E-04	1.04E-03
50/50	4	1	4	1.75E-01	1.53E-01	2.00E-01
50/50	4	1	1	1.68E-01	1.52E-01	2.09E-01
50/50	4	0.8	4	5.07E-01	3.87E-01	3.85E-01
50/50	4	0.8	1	4.67E-01	3.76E-01	4.13E-01
20/100	4	4	4	8.28E-04	8.27E-04	1.41E-03
20/100	4	4	1	8.29E-04	8.28E-04	1.42E-03
20/100	4	1	4	2.55E-01	1.97E-01	2.45E-01
20/100	4	1	1	2.48E-01	2.04E-01	2.60E-01
20/100	4	0.8	4	7.13E-01	4.15E-01	4.37E-01
20/100	4	0.8	1	6.66E-01	4.34E-01	4.78E-01

MSE in favor of rPCA



Simulation study

n/p	S	SNR	d_1/d_2	$MSE(\hat{\mathbf{X}}^{\text{PCA}}, \tilde{\mathbf{X}})$	$MSE(\hat{\mathbf{X}}^{\text{rPCA}}, \tilde{\mathbf{X}})$	$MSE(\hat{\mathbf{X}}^{\text{SURE}}, \tilde{\mathbf{X}})$
100/20	4	4	4	8.25E-04	8.24E-04	1.42E-03
100/20	4	4	1	8.26E-04	8.25E-04	1.43E-03
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50/50	4	4	1	5.46E-04	5.46E-04	1.04E-03
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20/100	4	0.8	4	7.13E-01	4.15E-01	4.37E-01
20/100	4	0.8	1	6.66E-01	4.34E-01	4.78E-01

rPCA $\geq$ PCA when SNR is small - rPCA $\approx$ PCA when SNR is high

Simulation study

n/p	S	SNR	d_1/d_2	$MSE(\hat{X}^{PCA}, \tilde{X})$	$MSE(\hat{X}^{rPCA}, \tilde{X})$	$MSE(\hat{X}^{SURE}, \tilde{X})$
100/20	4	4	4	8.25E-04	8.24E-04	1.42E-03
100/20	4	4	1	8.26E-04	8.25E-04	1.43E-03
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20/100	4	0.8	1	6.66E-01	4.34E-01	4.78E-01

$rPCA \geq PCA$ when $d_1 \geq d_2$



Simulation study

n/p	S	SNR	d_1/d_2	$MSE(\hat{\mathbf{X}}^{\text{PCA}}, \tilde{\mathbf{X}})$	$MSE(\hat{\mathbf{X}}^{\text{rPCA}}, \tilde{\mathbf{X}})$	$MSE(\hat{\mathbf{X}}^{\text{SURE}}, \tilde{\mathbf{X}})$
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20/100	4	0.8	1	6.66E-01	4.34E-01	4.78E-01

Same order of errors for $n/p = 0.2$ or 5, smaller for $n/p = 1$!



Simulation study

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100/20	4	1	1	2.47E-01	2.04E-01	2.62E-01
100/20	4	0.8	4	7.41E-01	4.27E-01	4.36E-01
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20/100	4	0.8	1	6.66E-01	4.34E-01	4.78E-01

SURE good when data are noisy



Simulation from Candes *et al.* (2012)

- Same model of simulation
- $n = 500$ $p = 200$, SNR (4, 2, 1, 0.5), S (10, 100)

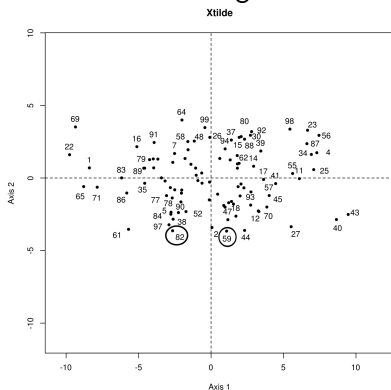
S	SNR	$MSE(\hat{\mathbf{X}}^{\text{PCA}}, \tilde{\mathbf{X}})$	$MSE(\hat{\mathbf{X}}^{\text{rPCA}}, \tilde{\mathbf{X}})$	$MSE(\hat{\mathbf{X}}^{\text{SURE}}, \tilde{\mathbf{X}})$
10	4	4.31E-03	4.29E-03	8.74E-03
10	2	1.74E-02	1.71E-02	3.29E-02
10	1	7.16E-02	6.75E-02	1.16E-01
10	0.5	3.19E-01	2.57E-01	3.53E-01
100	4	3.79E-02	3.69E-02	4.50E-02
100	2	1.58E-01	1.41E-01	1.56E-01
100	1	7.29E-01	4.91E-01	4.48E-01
100	0.5	3.16E+00	1.48E+00	8.52E-01



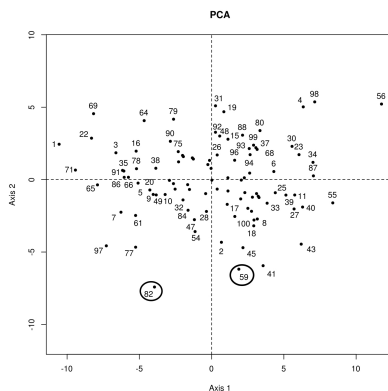
Individuals representation

100 individuals, 20 variables, 2 dimensions, $\text{SNR}=0.8$, $d_1/d_2 = 4$

True configuration



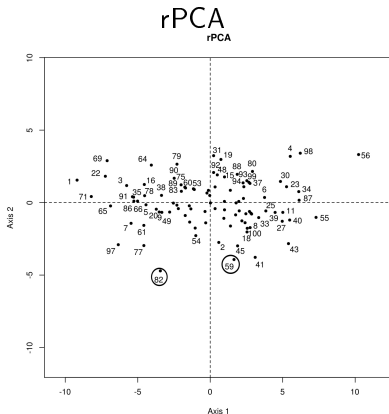
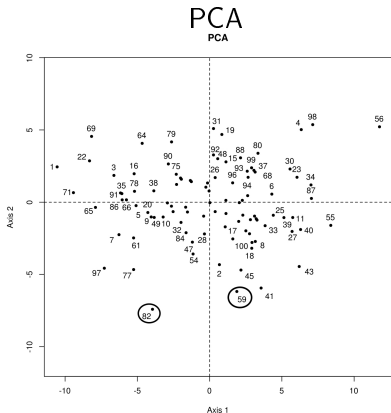
PCA





Individuals representation

100 individuals, 20 variables, 2 dimensions, $\text{SNR}=0.8$, $d_1/d_2 = 4$

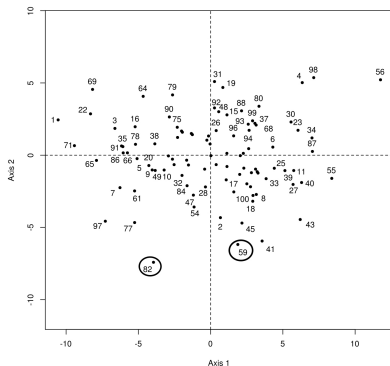




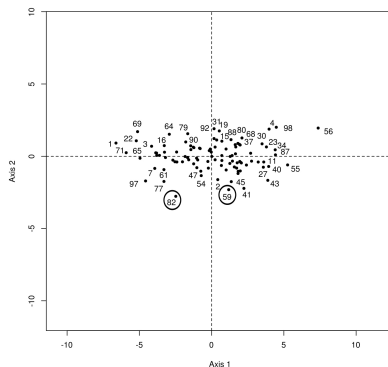
Individuals representation

100 individuals, 20 variables, 2 dimensions, $\text{SNR}=0.8$, $d_1/d_2 = 4$

PCA
PCA



SURE
SURE

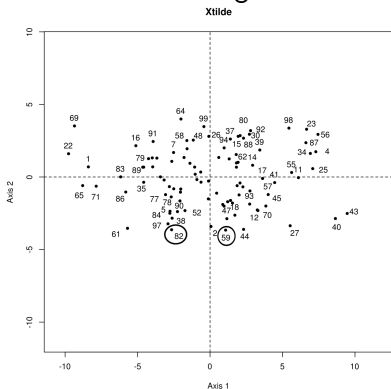




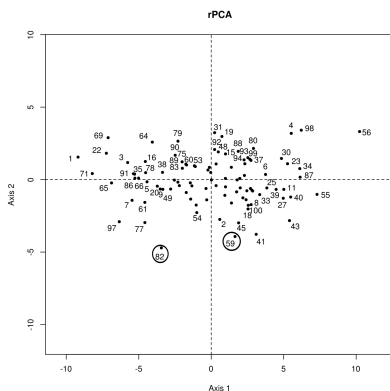
Individuals representation

100 individuals, 20 variables, 2 dimensions, $\text{SNR}=0.8$, $d_1/d_2 = 4$

True configuration



rPCA

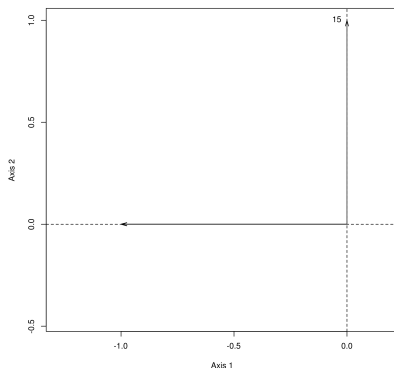


$\Rightarrow$ Improve the recovery of the inner-product matrix $\tilde{\mathbf{X}}\tilde{\mathbf{X}}'$

○○○○
○○○○

Variables representation

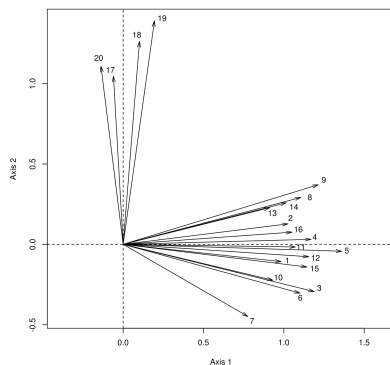
True configuration



$$\text{cor}(X_7, X_9) = 1$$

PCA

PCA

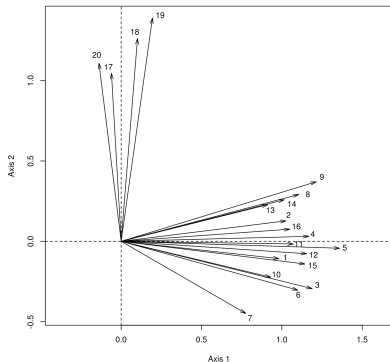


$$0.68$$



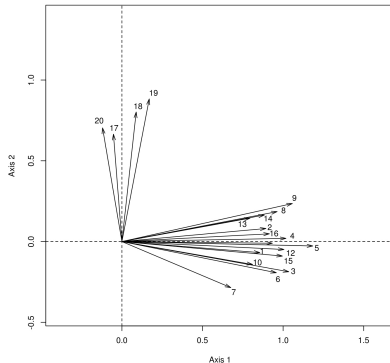
Variables representation

PCA
PCA



$$\text{cor}(X_7, X_9) = 0.68$$

rPCA
rPCA

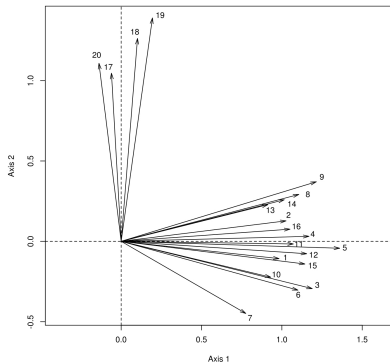


$$0.81$$



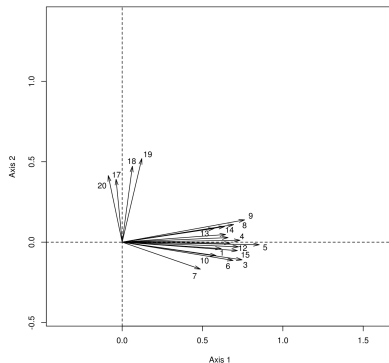
Variables representation

PCA
PCA



$$\text{cor}(X_7, X_9) = 0.68$$

SURE
SURE

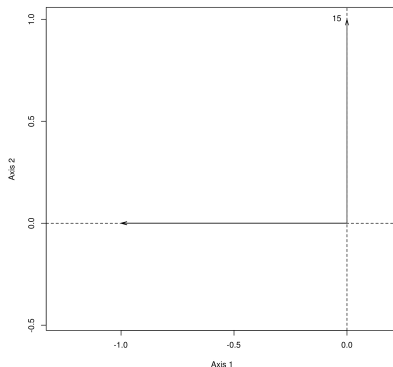


$$0.82$$

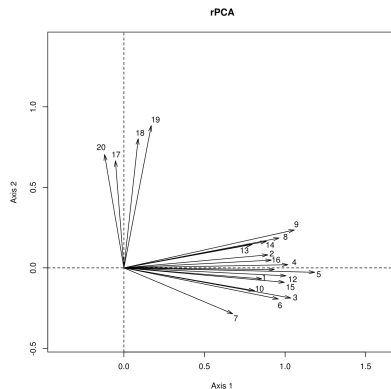


Variables representation

True configuration



rPCA



$$\text{cor}(X_7, X_9) = 1$$

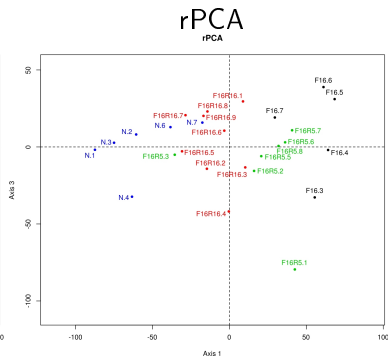
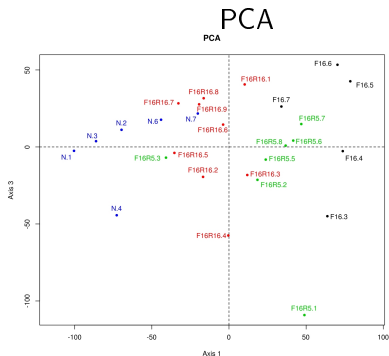
$$0.81$$

$\Rightarrow$ Improve the recovery of the covariance matrix $\tilde{\mathbf{X}}'\tilde{\mathbf{X}}$



Real data set

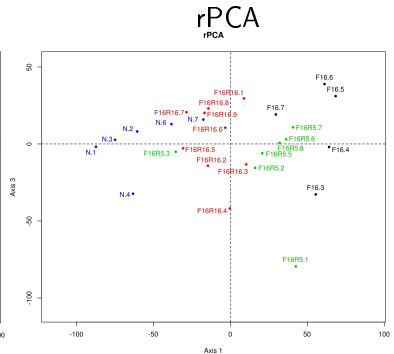
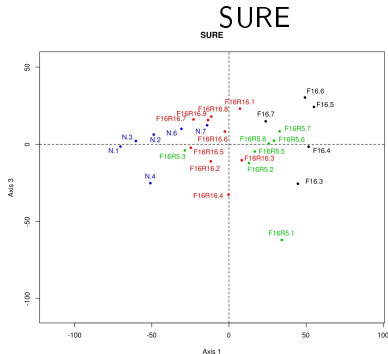
- 27 chickens submitted to 4 nutritional statuses
- 12664 gene expressions





Real data set

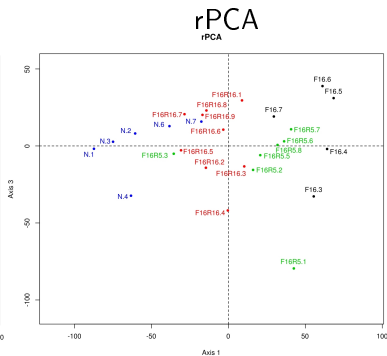
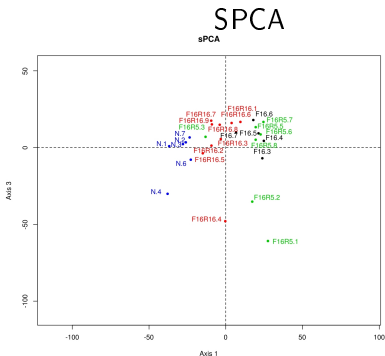
- 27 chickens submitted to 4 nutritional statuses
- 12664 gene expressions





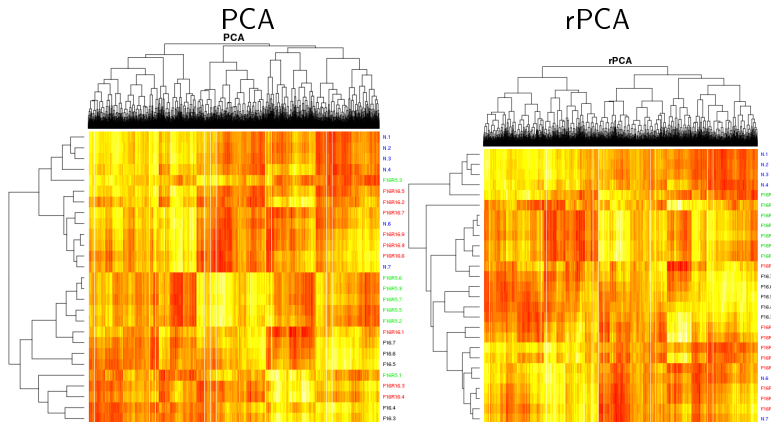
Real data set

- 27 chickens submitted to 4 nutritional statuses
- 12664 gene expressions



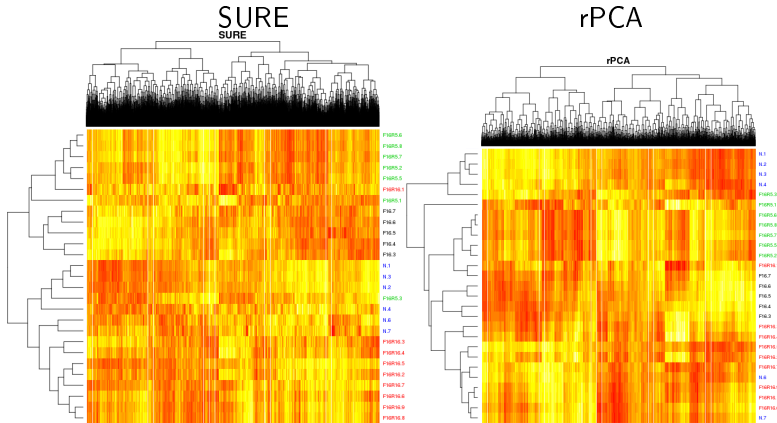


Real data set



Bi-clustering from rPCA better find the 4 nutritional statuses!

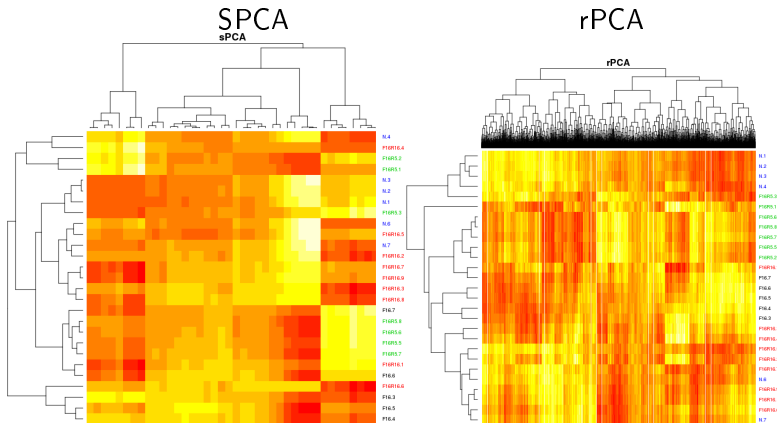
Real data set



Bi-clustering from rPCA better find the 4 nutritional statuses!

○○○○○
○○○○○

Real data set



Bi-clustering from rPCA better find the 4 nutritional statuses!



Conclusion

⇒ Recover a low rank matrix from noisy observations

- singular values thresholding: a popular estimation strategy
- a regularized term a priori: $\left(\sqrt{\lambda_s} - \frac{\hat{\sigma}^2}{\sqrt{\lambda_s}}\right)$: Minimize MSE - Bayesian fixed effect model

⇒ Good results in term of recovery of the structure, the distances between individuals and the covariance matrix and in term of visualization

- Asymptotic results
- Tuning parameter: the number of dimensions → CV?
- Regularized MCA (Takane, 2006)